

# More about extreme eigenvalues of perturbed random matrices

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# Outline of the talk

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- ▶ Large deviation principle

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$$\theta_1 \geq \dots \geq \theta_{r_0} > 0 > \theta_{r_0+1} \geq \dots \geq \theta_r.$$

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Main tool :

$$f_n(z) = \det \left( [G_{i,j}^n(z)]_{i,j=1}^r - \text{diag}(\theta_1^{-1}, \dots, \theta_r^{-1}) \right),$$

with

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Key point :

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We define

$$\rho_\theta := \begin{cases} G_{\mu_X}^{-1}(1/\theta) & \text{if } \theta \in (-\infty, \underline{\theta}) \cup (\bar{\theta}, +\infty), \\ a & \text{if } \theta \in [\underline{\theta}, 0) \\ b & \text{if } \theta \in (0, \bar{\theta}] \end{cases}$$

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and almost sure convergence of the extreme eigenvalues is governed by

## Theorem

For all  $i \in \{1, \dots, r_0\}$  we have

$$\tilde{\lambda}_i^n \xrightarrow{a.s.} \rho_{\theta_i}$$

and for all  $i > r_0$ ,

$$\tilde{\lambda}_i^n \xrightarrow{a.s.} b.$$

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$$\left( \gamma_i := \sqrt{n}(\tilde{\lambda}_i^n - \rho_{\theta_i}), i \in I_j \right)_{1 \leq j \leq q}$$

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We have to study, for  $\rho_n := \rho_\alpha + \frac{x}{\sqrt{n}}$ ,

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where

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- ▶ If the repulsion is milder, the extreme ev of  $\tilde{X}_n$  stick to the edge of the bulk.
- ▶ If the repulsion is even milder, the extreme ev of  $\tilde{X}_n$  stick to the extreme ev of  $X_n$  even at the level of fluctuations.

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$$\sum_{i=m_n+1}^n \frac{1}{(\lambda_r - \lambda_i)^2} \leq n^{2-\eta}, \quad \sum_{i=m_n+1}^n \frac{1}{(\lambda_r - \lambda_i)^4} \leq n^{4-\eta'}$$

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then if  $I_b$  is the set of indices for which  $\rho_{\theta_i} = b$ , for any  $\alpha' > \alpha$ , with overwhelming probability,

$$\max_{i \in I_b} \min_k |\tilde{\lambda}_i - \lambda_k| \leq n^{-1+\alpha'}.$$

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so that

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Open question : fluctuations for critical  $\theta_i$ 's.

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In the iid case,  $f_n$  depends polynomially on the entries of  $K^n(z)$  with

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We can find  $H^n$  having the same zeroes as  $f_n$  and depending polynomially on the entries of  $K^n(z)$  and  $C^n$

$$(K^n(z))_{ij} := \frac{1}{n} \sum_{k=1}^n \frac{g_i(k)g_j(k)}{z - \lambda_k} \quad \text{and} \quad (C^n)_{ij} := \frac{1}{n} \sum_{k=1}^n g_i(k)g_j(k).$$

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$X_n$  diagonal, deterministic, satisfying (H1).

$G = (g_1, \dots, g_r)$  a random vector satisfying that  $\mathbb{E}(e^{\alpha \sum |g_i^2|}) < \infty$  for some  $\alpha > 0$  (and not charging an hyperplane)

$G_i^n$  random vector whose entries are  $1/\sqrt{n}$  times independent copies of  $g_i$  and  $U_i^n$  obtained by orthonormalization.

We can find  $H^n$  having the same zeroes as  $f_n$  and depending polynomially on the entries of  $K^n(z)$  and  $C^n$

$$(K^n(z))_{ij} := \frac{1}{n} \sum_{k=1}^n \frac{g_i(k)g_j(k)}{z - \lambda_k} \quad \text{and} \quad (C^n)_{ij} := \frac{1}{n} \sum_{k=1}^n g_i(k)g_j(k).$$

## Theorem

The law of the  $r_0$  largest eigenvalues of  $\widetilde{X}_n$  satisfies a LDP in the scale  $n$  with a good rate function.

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Remark : minimizers depend on  $G$  only through its covariance matrix.

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$$\mathbf{I}(K(\cdot), C) = \sup_{P, X, Y} \left\{ \text{Tr} \left( \int K'(z) P(z) dz + K(z^*) X + C Y \right) - \Gamma(P, Y, X) \right\}$$

where  $\Gamma(P, Y, X)$  is given by the formula

$$\Gamma(P, Y, X) = \int \Lambda \left( - \int \frac{1}{(z-x)^2} P(z) dz + \frac{1}{z^* - x} X + Y \right) d\mu_X(x)$$

and the supremum is taken over piecewise constant functions  $P$  with values in  $H_r$  and  $X, Y$  in  $H_r$ .

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$$J_{\mathcal{K}}(f) = \inf\{\mathbf{I}(F) : F \in \mathcal{C}(\mathcal{K}, H_r) \times H_r, P_{\Theta}(F(z)) = f(z), \forall z \in \mathcal{K}\}.$$

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### Theorem

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$$L(\alpha) = \begin{cases} \lim_{\varepsilon \downarrow 0} \inf_{\cup_{\gamma > 0} S_{(\alpha_1, \dots, \alpha_{m-k}), \gamma}^{\varepsilon}} J_{\mathcal{K}_{\varepsilon}} & \text{if } \alpha \in \mathbb{R}_{\downarrow}^m(b), \alpha_{m-k+1} = b \text{ and} \\ & \alpha_{m-k} > b, \\ +\infty & \text{otherwise.} \end{cases}$$

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with

$$S_{\alpha, \gamma}^{\varepsilon} := \left\{ f \in \mathcal{C}(K_{\varepsilon}) : f(z) = s \cdot g(z) \prod_{i=1}^{m-k} (z - \alpha_i) \text{ with } g \geq \gamma \right\},$$

# Study of the minimizers

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$L$  good rate function, vanishes at minimizers  $(\lambda_1^*, \dots, \lambda_m^*)$ . Let  $k$  be such that  $\lambda_{m-k}^* > b$  and  $\lambda_{m-k+1}^* = b$ . By compactity, one can find  $f$  vanishing at  $(\lambda_1^*, \dots, \lambda_{m-k}^*)$  such that  $J_{K_\varepsilon}(f) = 0$  for any  $\varepsilon > 0$ . It also means that  $f(z) = P_\Theta(K, C)$ , with  $(K, C)$  minimizing **I**.



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$$\left| \mathbb{E} \left( e^{\varepsilon \text{Tr} \left( - \int \frac{1}{(z-x)^2} P(z) z + \frac{1}{z^* - x} X + Y \right) Z} \right) \right. \\ \left. - \mathbb{E} \left( 1 + \varepsilon \text{Tr} \left( \left( - \int \frac{1}{(z-x)^2} P(z) dz + \frac{1}{z^* - x} X + Y \right) Z \right) \right) \right| \leq \varepsilon^2 L,$$

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$$\left| \mathbb{E} \left( e^{\varepsilon \text{Tr} \left( - \int \frac{1}{(z-x)^2} P(z) z + \frac{1}{z^* - x} X + Y \right) Z} \right) - \mathbb{E} \left( 1 + \varepsilon \text{Tr} \left( \left( - \int \frac{1}{(z-x)^2} P(z) dz + \frac{1}{z^* - x} X + Y \right) Z \right) \right) \right| \leq \varepsilon^2 L,$$

so that

$$\Gamma(\varepsilon P, \varepsilon X, \varepsilon Y) = \varepsilon \text{Tr} \left( \int (K^*)'(z) P(z) dz + K^*(z^*) X + C^* Y \right) + O(\varepsilon^2)$$

with

$$(K^*(z))_{ij} = \int \frac{(C^*)_{ij}}{z - \lambda} d\mu_X(\lambda) \quad \text{and} \quad (C^*)_{ij} = \mathbb{E}[g_i g_j].$$

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In the case when  $(g_1, \dots, g_r)$  are independent centered variables with variance one, one can check that  $C^* = I_r$ ,  $K^*(z) = \int \frac{1}{z-x} \mu_X(x) \cdot I_r$  and

$$H(z) = \prod_{i=1}^r \left( \frac{1}{\theta_i} - \int \frac{1}{z-x} \mu_X(x) \right)$$